

## FLIPS IN GRAPHS\*

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**Abstract.** We study a problem motivated by a question related to quantum error-correcting codes. Combinatorially, it involves the graph parameter  $f(G) = \min\{|A| + |\{x \in V \setminus A : d_A(x) \text{ is odd}\}| : A \neq \emptyset\}$ , where  $V$  is the vertex set of  $G$  and  $d_A(x)$  is the number of neighbors of  $x$  in  $A$ . We give asymptotically tight estimates of  $f$  for the random graph  $G_{n,p}$  when  $p$  is constant. Also, if  $f(n) = \max\{f(G) : |V(G)| = n\}$ , then we show that  $f(n) \leq (0.382 + o(1))n$ .

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**1. Introduction.** In this paper we consider a problem which is motivated by a question from quantum error-correcting codes.

Given a graph  $G$  with  $\pm 1$  signs on vertices, each vertex can perform at most one of the following three operations:  $O_1$  (flip all neighbors, i.e., change their signs),  $O_2$  (flip oneself), and  $O_3$  (flip oneself and all neighbors). We want to start with all  $+1$ 's, execute some nonzero number of operations, and return to all  $+1$ 's. The *diagonal distance*  $f(G)$  is the minimum number of operations needed (with each vertex doing at most one operation).

Trivially,

$$(1.1) \quad f(G) \leq \delta(G) + 1$$

holds, where  $\delta(G)$  denotes the minimum degree. Indeed, a vertex with the minimum degree applies  $O_1$  and then its neighbors fix themselves applying  $O_2$ . Let

$$f(n) = \max f(G),$$

where the maximum is taken over all nonempty graphs of order  $n$ .

Given a graph  $G$ , one can ultimately construct a quantum error-correcting code; see [3, 5, 6]. A common metric to measure the code robustness against noise is the quantity called “code distance” which is bounded from above by  $f(G)$ . Although it is more important to find explicit graphs  $G$  with large  $f(G)$  (see the case  $k = 0$  of section “QECC” in [2] for known constructions), theoretical upper and lower bounds on  $f(n)$  are also of interest.

In this paper we asymptotically determine the diagonal distance of the random graph  $G_{n,p}$  for any  $p \in (0, 1)$ .

We denote the *symmetric difference* of two sets  $A$  and  $B$  by  $A \triangle B$  and the *logarithmic function* with base  $e$  as  $\log$ .

**THEOREM 1.1.** *There are absolute constants  $\lambda_0 \approx 0.189$  and  $p_0 \approx 0.894$  (see (2.4) and (3.3)) such that for  $G = G_{n,p}$  asymptotically almost surely*

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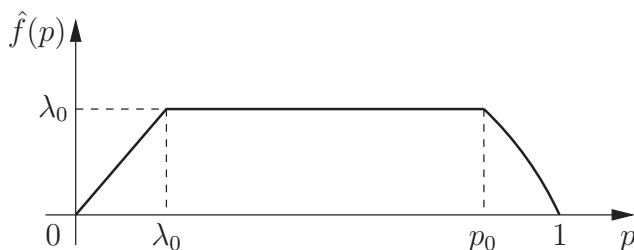


FIG. 1.1. The behavior of  $\hat{f}(p) = \lim_{n \rightarrow \infty} f(G_{n,p})/n$  as a function of  $p$ .

- (i)  $f(G) = \delta(G) + 1$  for constant  $0 < p < \lambda_0$  or  $p = o(1)$ ,
- (ii)  $|f(G) - \lambda_0 n| = \tilde{O}(n^{1/2})$  for  $\lambda_0 \leq p \leq p_0$ ,
- (iii)  $f(G) = 2 + \min_{x,y \in V(G)} |(N(x) \triangle N(y)) \setminus \{x,y\}|$  for constant  $p_0 < p < 1$  or  $p = 1 - o(1)$ .

(Here  $\tilde{O}(n^{1/2})$  hides a polylog factor.)

Figure 1.1 visualizes the behavior of the diagonal distance of  $G_{n,p}$ . In addition to Theorem 1.1 we find the following upper bound on  $f(n)$ .

**THEOREM 1.2.**  $f(n) \leq (0.382 + o(1))n$ .

In the remainder of the paper we will use a more convenient restatement of  $f(G)$ . Observe that the order of execution of operations does not affect the final outcome. For any  $A \subset V = V(G)$ , let  $B$  consist of those vertices in  $V \setminus A$  that have an odd number of neighbors in  $A$ . Let  $a = |A|$  and  $b = |B|$ . Then  $f(G)$  is the minimum of  $a + b$  over all nonempty  $A \subset V(G)$ . The vertices of  $A$  do an  $O_1/O_3$  operation, depending on the even/odd parity of their neighborhood in  $A$ . The vertices in  $B$  then do an  $O_2$ -operation to change back to  $+1$ .

**2. Random graphs for  $p = 1/2$ .** Here we prove a special case of Theorem 1.1 when  $p = 1/2$ . This case is somewhat easier to handle.

Let  $G = G_{n,1/2}$  be a binomial random graph. First we find a lower bound on  $f(G)$ . If we choose a nonempty  $A \subset V$  and then generate  $G$ , then the distribution of  $b$  is binomial with parameters  $n - a$  and  $1/2$ , which we denote here by  $Bin(n - a, 1/2)$ . Hence, if  $l$  is such that

$$(2.1) \quad \sum_{a=1}^{l-1} \binom{n}{a} \Pr(Bin(n - a, 1/2) \leq l - 1 - a) = o(1),$$

then asymptotically almost surely the diagonal distance of  $G$  is at least  $l$ .

Let  $\lambda = l/n$  and  $\alpha = a/n$ . We may assume that  $\lambda < \frac{1}{2}$ . Consequently,  $\lambda - \alpha < \frac{1}{2}(1 - \alpha)$ , and hence we can approximate the summand in (2.1) by

$$2^{n(H(\alpha) + (1-\alpha)(H(\lambda - \alpha / (1-\alpha)) - 1) + O(\log n/n))},$$

where  $H$  is the binary entropy function defined as  $H(p) = -p \log_2 p - (1-p) \log_2 (1-p)$ . For more information about the entropy function and its properties see, e.g., [1]. Let

$$(2.2) \quad g_\lambda(\alpha) = H(\alpha) + (1 - \alpha) \left( H \left( \frac{\lambda - \alpha}{1 - \alpha} \right) - 1 \right).$$

The maximum of  $g_\lambda(\alpha)$  is attained exactly for  $\alpha = 2\lambda/3$ , since

$$g'_\lambda(\alpha) = \log_2 \frac{2(\lambda - \alpha)}{\alpha}.$$

Now the function

$$(2.3) \quad h(\lambda) = g_\lambda(2\lambda/3)$$

is concave on  $\lambda \in [0, 1]$  since

$$h''(\lambda) = \frac{1}{(\lambda - 1)\lambda \log 2} < 0.$$

Moreover, observe that  $h(0) = -1$  and  $h(1) = H(2/3) - 1/3 > 0$ . Thus the equation  $h(\lambda) = 0$  has a unique solution  $\lambda_0$ , and one can compute that

$$(2.4) \quad \lambda_0 = 0.1892896249152306 \dots$$

Therefore, if  $\lambda = \lambda_0 - K \log n/n$  for large enough  $K > 0$ , then the left-hand side of (2.1) goes to zero and similarly for  $\lambda = \lambda_0 + K \log n/n$  it goes to infinity. In particular,  $f(G) > (\lambda_0 - o(1))n$  asymptotically almost surely.

Let us show that this constant  $\lambda_0$  is best possible, i.e., asymptotically almost surely  $f(G) \leq (\lambda_0 + K \log n/n)n$ . Let  $\lambda = \lambda_0 + K \log n/n$ ,  $n$  be large, and  $l = \lambda n$ . Let  $\alpha = 2\lambda/3$ , and  $a = \lfloor \alpha n \rfloor$ . We pick a random  $a$ -set  $A \subset V$  and compute  $b$ . Let  $X_A$  be an indicator random variable so that  $X_A = 1$  if and only if  $b = b(A) \leq l - a$ . Let  $X = \sum_{|A|=a} X_A$ . We succeed if  $X > 0$ .

The expectation  $E(X) = \binom{n}{a} \Pr(\text{Bin}(n - a, 1/2) \leq l - a)$  tends to infinity by our choice of  $\lambda$ . We now show that  $X > 0$  asymptotically almost surely by using the Chebyshev inequality. First note that for  $A \cap C \neq \emptyset$  we have

$$\text{Cov}(X_A, X_C) = \Pr(X_A = X_C = 1) - \Pr(X_A = 1)\Pr(X_C = 1) = 0.$$

Indeed, if  $x \in V \setminus (A \cup C)$ , then  $\Pr(x \in B(A) | X_C = 1) = 1/2$ , since  $A \setminus C \neq \emptyset$  and no adjacency between  $x$  and all vertices in  $A \setminus C$  is exposed by the event  $X_C = 1$ . Similarly, if  $x \in C \setminus A$ , then  $A \cap C \neq \emptyset$  and an adjacency between  $x$  and  $A \cap C$  is independent of the occurrence of  $X_C = 1$ . This implies that  $\Pr(x \in B(A) | X_C = 1) = 1/2$  as well. Thus  $\Pr(X_A = 1 | X_C = 1) = \Pr(\text{Bin}(n - a, 1/2) \leq l - a) = \Pr(X_A = 1)$ , and consequently  $\text{Cov}(X_A, X_C) = 0$ .

Now consider the case when  $A \cap C = \emptyset$ . Let  $s$  be a vertex in  $A$ . Define a new indicator random variable  $Y$  which takes the value 1 if and only if  $|B(C) \setminus \{s\}| \leq l - a$ . Observe that

$$\begin{aligned} \Pr(Y = 1) &= \Pr(\text{Bin}(n - a - 1, 1/2) \leq l - a) \\ &\leq 2 \Pr(\text{Bin}(n - a, 1/2) \leq l - a) = 2 \Pr(X_A = 1). \end{aligned}$$

Moreover,

$$\Pr(X_A = 1 | Y = 1) = \Pr(\text{Bin}(n - a, 1/2) \leq l - a) = \Pr(X_A = 1),$$

since for every  $x \in V \setminus A$  the adjacency between  $x$  and  $s$  is not influenced by  $Y = 1$ . Finally note that  $X_C \leq Y$ . Thus,

$$\begin{aligned} \text{Cov}(X_A, X_C) &\leq \Pr(X_A = X_C = 1) \\ &\leq \Pr(X_A = Y = 1) = \Pr(Y = 1)\Pr(X_A = 1 | Y = 1) \leq 2(\Pr(X_A = 1))^2. \end{aligned}$$

Consequently,

$$\begin{aligned} \text{Var}(X) &= E(X) + \sum_{A \cap C \neq \emptyset, A \neq C} \text{Cov}(X_A, X_C) + \sum_{A \cap C = \emptyset} \text{Cov}(X_A, X_C) \\ &\leq E(X) + 2 \sum_{A \cap C = \emptyset} (\Pr(X_A = 1))^2 \\ &= E(X) + 2 \binom{n}{a} \binom{n-a}{a} (\Pr(X_A = 1))^2 = o(E(X)^2), \end{aligned}$$

as  $E(X) = \binom{n}{a} \Pr(X_A = 1)$  tends to infinity and  $\binom{n-a}{a} = o(\binom{n}{a})$ . Hence, Chebyshev’s inequality yields that  $X > 0$  asymptotically almost surely.

*Remark 2.1.* A version of the well-known Gilbert–Varshamov bound (see, e.g., [4]) states that if

$$(2.5) \quad 2^{-n} \sum_{i=1}^{l-1} \binom{n}{i} 3^i < 1,$$

then  $f(n) \geq l$ . Observe that this is consistent with bound (2.1). Let  $\lambda = l/n$ . We can approximate the left-hand side of (2.5) by

$$2^{n(H(\lambda) + \lambda \log_2 3 - 1 + o(1))}.$$

One can check after some computation that

$$H(\lambda) + \lambda \log_2 3 - 1 = g_\lambda(2\lambda/3).$$

Therefore, (2.1) and (2.5) give asymptotically the same lower bound on  $f(n)$ .

**3. Random graphs for arbitrary  $p$ .** Let  $G = G_{n,p}$  be a random graph with  $p \in (0, 1)$ .

Observe that for a fixed set  $A \subset V$ ,  $|A| = a$ , the probability that a vertex from  $V \setminus A$  belongs to  $B(A)$  is

$$p(a) = \sum_{0 \leq i < \frac{a}{2}} \binom{a}{2i+1} p^{2i+1} (1-p)^{a-(2i+1)} = \frac{1 - (1-2p)^a}{2}.$$

(If this is unfamiliar, write  $1 - (1-2p)^a = ((1-p) + p)^a - ((1-p) - p)^a$  and expand.)

**3.1.  $0 < p < \lambda_0$ .** For  $p < \lambda_0$  we begin with the upper bound  $f(G) \leq \delta(G) + 1$ ; see (1.1). For the lower bound it is enough to show that

$$(3.1) \quad \sum_{2 \leq a \leq pn} \binom{n}{a} \Pr(\text{Bin}(n-a, p(a)) \leq pn-a) = o(1),$$

since  $\delta(G) + 1 \leq np$  asymptotically almost surely. (We may assume that  $p = \Omega(\frac{\log n}{n})$ ; otherwise  $\delta(G) = 0$  with high probability, and the theorem is trivially true.) This implies that with high probability if  $|A| + |B| \leq pn$ , then  $|A| = 1$ .

**3.1.1.  $p$  constant.** We split this sum into two sums for  $2 \leq a \leq \sqrt{n}$  and  $\sqrt{n} < a \leq pn$ , respectively. Let  $X = \text{Bin}(n-a, p(a))$  and

$$\varepsilon = 1 - \frac{pn-a}{(n-a)p(a)} \geq 1 - \frac{p}{p(2)} = 1 - \frac{1}{2-2p} > 0.$$

We will use the following version of Chernoff’s bound:

$$\Pr(\text{Bin}(N, \rho) \leq (1 - \theta)N\rho) \leq e^{-\theta^2 N\rho/2}.$$

Hence, we see that

$$\begin{aligned} \Pr(\text{Bin}(n - a, p(a)) \leq pn - a) \\ = \Pr(X \leq (1 - \varepsilon)E(X)) \leq \exp\{-\varepsilon^2 E(X)/2\} = \exp\{-\Theta(n)\}, \end{aligned}$$

and consequently

$$\begin{aligned} \sum_{2 \leq a < \sqrt{n}} \binom{n}{a} \Pr(\text{Bin}(n - a, p(a)) \leq pn - a) \\ \leq \sqrt{n} \binom{n}{\sqrt{n}} \exp\{-\Theta(n)\} \leq \exp\{O(\sqrt{n} \log n)\} \exp\{-\Theta(n)\} = o(1). \end{aligned}$$

Now we bound the second sum corresponding to  $\sqrt{n} < a \leq pn$ . Note that

$$\begin{aligned} \sum_{\sqrt{n} \leq a \leq pn} \binom{n}{a} \Pr(\text{Bin}(n - a, p(a)) \leq pn - a) \\ = \sum_{\sqrt{n} \leq a \leq pn} \binom{n}{a} \Pr\left(\text{Bin}\left(n - a, \frac{1}{2} + e^{-\Omega(n^{1/2})}\right) \leq pn - a\right) \\ \leq n2^{n(h(p)+o(1))} = o(1). \end{aligned}$$

Here  $h$  is defined in (2.3) and the right-hand limit is zero since  $p < \lambda_0$ .

**3.1.2.  $p = o(1)$ .** We follow basically the same strategy as above and show that (3.1) holds for large  $a$  and something similar when  $a$  is small. Suppose then that  $p = 1/\omega$ , where  $\omega = \omega(n) \rightarrow \infty$ . First consider those  $a$  for which  $ap \geq 1/\omega^{1/2}$ . In this case  $p(a) \geq (1 - e^{-2ap})/2$ . Thus,

$$\begin{aligned} \sum_{\substack{ap \geq 1/\omega^{1/2} \\ a \leq np}} \binom{n}{a} \Pr(\text{Bin}(n - a, p(a)) \leq pn - a) \\ = \sum_{\substack{ap \geq 1/\omega^{1/2} \\ a \leq np}} e^{O(n \log \omega/\omega)} e^{-\Omega(n/\omega^{1/2})} = o(1). \end{aligned}$$

If  $ap \leq 1/\omega^{1/2}$ , then  $p(a) = ap(1 + O(ap))$ . Then

$$(3.2) \quad \sum_{\substack{ap < 1/\omega^{1/2} \\ 2 \leq a \leq np}} \binom{n}{a} \Pr(\text{Bin}(n - a, p(a)) \leq pn - a) \leq \sum_{\substack{ap < 1/\omega^{1/2} \\ 2 \leq a \leq np}} \left(\frac{ne}{a} e^{-np/10}\right)^a = o(1),$$

provided  $np \geq 11 \log n$ .

If  $np \leq \log n - \log \log n$ , then  $G = G_{n,p}$  has isolated vertices asymptotically almost surely and then  $f(G) = 1$ . So we are left with the case where  $\log n - \log \log n \leq np \leq 11 \log n$ .

We next observe that if there is a set  $A$  for which  $2 \leq |A|$  and  $|A| + |B(A)| \leq np$ , then there is a minimal size such set. Let  $H_A = (A, E_A)$  be a graph with vertex set  $A$  and an edge  $(v, w) \in E_A$  if and only if  $v, w$  have a common neighbor in  $G$ .  $H_A$  must be connected, else  $A$  is not minimal. So we can find  $t \leq a - 1$  vertices  $T$  such that  $A \cup T$  spans at least  $t + a - 1$  edges between  $A$  and  $T$ . Thus we can replace the estimate (3.2) by

$$\begin{aligned} & \sum_{\substack{ap < 1/\omega^{1/2} \\ 2 \leq a \leq np}} \sum_{t=1}^{a-1} \binom{n}{a} \binom{n}{t} \binom{ta}{t+a-1} p^{t+a-1} \Pr(\text{Bin}(n-a-t, p(a)) \leq pn-a) \\ & \leq \sum_{\substack{ap < 1/\omega^{1/2} \\ 2 \leq a \leq np}} \sum_{t=1}^{a-1} \left(\frac{ne}{a}\right)^a \left(\frac{ne}{t}\right)^t \left(\frac{taep}{t+a-1}\right)^{t+a-1} e^{-anp/10} \\ & \leq \frac{1}{e^2 np} \sum_{\substack{ap < 1/\omega^{1/2} \\ 2 \leq a \leq np}} a \left((e^2 np)^2 e^{-np/10}\right)^a = o(1). \end{aligned}$$

**3.2.  $p_0 < p < 1$ .** First let us define the constant  $p_0$ . Let

$$(3.3) \quad p_0 \approx 0.8941512242051071 \dots$$

be a root of  $2p - 2p^2 = \lambda_0$ . For the upper bound let  $A = \{x, y\}$ , where  $x$  and  $y$  satisfy  $|N(x) \Delta N(y)| \leq |N(x') \Delta N(y')|$  for any  $x', y' \in V(G)$ . Then  $B = B(A) = N(x) \Delta N(y)$ , and thus asymptotically almost surely  $|B| \leq (2p - 2p^2)n$  plus a negligible error term  $o(n)$ . (We may assume that  $1 - p = \Omega(\frac{\log n}{n})$ ; otherwise we have two vertices of degree  $n - 1$  with high probability, and hence  $f(G) = 2$ .)

To show the lower bound it is enough to prove that

$$\sum_{3 \leq a \leq (2p-2p^2)n} \binom{n}{a} \Pr(\text{Bin}(n-a, p(a)) \leq (2p-2p^2)n-a) = o(1).$$

Indeed, this implies that if  $|A| + |B| \leq (2p - 2p^2)n$ , then  $|A| = 1$  or  $2$ . But if  $|A| = 1$ , then in a typical graph  $|B| = (p + o(1))n > (2p - 2p^2)n$  since  $p > 1/2$ .

**3.2.1.  $p$  constant.** As in the previous section, we split the sum into two sums for  $3 \leq a \leq \sqrt{n}$  and  $\sqrt{n} < a \leq pn$ , respectively. Let

$$\varepsilon = 1 - \frac{(2p - 2p^2)n - a}{(n - a)p(a)} \geq 1 - \frac{2p - 2p^2}{p(a)} > 0.$$

To confirm the second inequality we have to consider two cases. The first one is for  $a$  odd and at least 3. Here,

$$1 - \frac{2p - 2p^2}{p(a)} > 1 - \frac{2p - 2p^2}{1/2} = (2p - 1)^2 > 0.$$

The second case, for  $a$  even and at least 4, gives

$$1 - \frac{2p - 2p^2}{p(a)} > 1 - \frac{2p - 2p^2}{p(2)} = 0.$$

Now one can apply Chernoff bounds with the given  $\varepsilon$  to show that

$$\sum_{3 \leq a < \sqrt{n}} \binom{n}{a} \Pr(Bin(n - a, p(a)) \leq (2p - 2p^2)n - a) = o(1).$$

Now we bound the second sum corresponding to  $\sqrt{n} < a \leq (2p - 2p^2)n$ . Note that

$$\begin{aligned} & \sum_{\sqrt{n} \leq a \leq (2p-2p^2)n} \binom{n}{a} \Pr(Bin(n - a, p(a)) \leq (2p - 2p^2)n - a) \\ &= \sum_{\sqrt{n} \leq a \leq (2p-2p^2)n} \binom{n}{a} \Pr\left(Bin\left(n - a, \frac{1}{2} + O(e^{-\Omega(n^{1/2})})\right) \leq (2p - 2p^2)n - a\right) \\ & \leq n2^{nh(2p-2p^2)+o(1)} = o(1) \end{aligned}$$

since  $p > p_0$  implies that  $2p - 2p^2 < \lambda_0$ .

**3.2.2.  $p = 1 - o(1)$ .** One can check it by following the same strategy as above and in section 3.1.2.

**3.3.  $\lambda_0 \leq p \leq p_0$ .** Let  $\alpha = 2\lambda_0/3$ ,  $a = \lfloor \alpha n \rfloor$ . Fix an  $a$ -set  $A \subset V$ , generate our random graph, and determine  $B = B(A)$  with  $b = |B|$ . Let  $\varepsilon = (\log n)^4/\sqrt{n}$ , and let  $X_A$  be the indicator random variable for  $a + b \leq (\lambda_0 + \varepsilon)n$  and  $X = \sum_A X_A$ . Then

$$p(a) = \frac{1}{2} + e^{-\Omega(n)},$$

and with  $g_\lambda(\alpha)$  as defined in (2.2),

$$(3.4) \quad E(X) = \exp\{(g_{\lambda_0+\varepsilon}(2\lambda_0/3) + o(1))n \log 2\}.$$

Now

$$\begin{aligned} g_{\lambda+\varepsilon}(\alpha) &= g_\lambda(\alpha) + (1 - \alpha) \left( H\left(\frac{\lambda + \varepsilon - \alpha}{1 - \alpha}\right) - H\left(\frac{\lambda - \alpha}{1 - \alpha}\right) \right) \\ &= g_\lambda(\alpha) + \varepsilon \log_2\left(\frac{1 - \lambda}{\lambda - \alpha}\right) + O(\varepsilon^2). \end{aligned}$$

Plugging this into (3.4) with  $\lambda = \lambda_0$  and  $\alpha = 2\lambda_0/3$  we see that

$$(3.5) \quad E(X) = \exp\left\{\left(\varepsilon \log_2\left(\frac{1 - \lambda_0}{\lambda_0/3}\right) + O(\varepsilon^2)\right) n \log 2\right\} = e^{\Omega((\log n)^4 n^{1/2})}.$$

Next, we estimate the variance of  $X$ . We will argue that for  $A, C \in \binom{V}{a}$  either  $|A \Delta C|$  is small (but the number of such pairs is small) or  $|A \Delta C|$  is large (but then the covariance  $Cov(X_A, X_C)$  is very small since if we fix the adjacency of some vertex  $x$  to  $C$ , then the parity of  $|N(x) \cap (A \setminus C)|$  is almost a fair coin flip). Formally,

$$\begin{aligned} Var(X) &= E(X) + \sum_{A \neq C} Cov(X_A, X_C) \\ &\leq E(X) + \sum_{|A \Delta C| < 2\sqrt{n}} \Pr(X_A = X_C = 1) \\ &\quad + \sum_{|A \Delta C| \geq 2\sqrt{n}, |A \cap C| \geq \sqrt{n}} Cov(X_A, X_C) \\ &\quad + \sum_{|A \cap C| < \sqrt{n}} \Pr(X_A = X_C = 1). \end{aligned}$$

Since  $E(X)$  goes to infinity, clearly  $E(X) = o(E(X)^2)$ . We show in Claims 3.1, 3.2, and 3.3 that the remaining part is also bounded by  $o(E(X)^2)$ . Then Chebyshev's inequality will imply that  $X > 0$  asymptotically almost surely.

CLAIM 3.1.  $\sum_{|A \Delta C| < 2\sqrt{n}} \Pr(X_A = X_C = 1) = o(E(X)^2)$

*Proof.* We estimate trivially  $\Pr(X_A = X_C = 1) \leq \Pr(X_A = 1)$ . Then

$$\begin{aligned} \sum_{|A \Delta C| < 2\sqrt{n}} \Pr(X_A = 1) &= \binom{n}{a} \sum_{0 \leq i < \sqrt{n}} \binom{n-a}{i} \binom{a}{a-i} \Pr(X_A = 1) \\ &= E(X) \sum_{0 \leq i < \sqrt{n}} \binom{n-a}{i} \binom{a}{a-i} \leq E(X) 2^{O(\sqrt{n} \log n)}. \end{aligned}$$

Thus, (3.5) yields that  $\sum_{|A \Delta C| < 2\sqrt{n}} \Pr(X_A = X_C = 1) = o(E(X)^2)$ . □

CLAIM 3.2.  $\sum_{|A \Delta C| \geq 2\sqrt{n}, |A \cap C| \geq \sqrt{n}} Cov(X_A, X_C) = o(E(X)^2)$ .

*Proof.* If  $x \in V \setminus (A \cup C)$ , then  $\Pr(x \in B(A) | X_C = 1) = 2^{-1+o(1/n)}$ , since we can always find at least  $\sqrt{n}$  vertices in  $A \setminus C$  with no adjacency with  $x$  determined by the event  $X_C = 1$ . Similarly, if  $x \in C \setminus A$ , then there are at least  $\sqrt{n} - 1$  vertices in  $A \cap C$  such that their adjacency with  $x$  is independent of the occurrence of  $X_C = 1$ . This implies that

$$\Pr(X_A = 1 | X_C = 1) = \sum_{0 \leq i \leq t-a} \binom{n-a}{i} 2^{-(n-a)+o(1)} = 2^{o(1)} \Pr(X_A = 1),$$

and consequently  $Cov(X_A, X_C) = o(\Pr(X_A = 1)^2)$ . Hence,

$$\sum_{|A \Delta C| \geq 2\sqrt{n}, |A \cap C| \geq \sqrt{n}} Cov(X_A, X_C) \leq \binom{n}{a}^2 o(\Pr(X_A = 1)^2) = o(E(X)^2). \quad \square$$

CLAIM 3.3.  $\sum_{|A \cap C| < \sqrt{n}} \Pr(X_A = X_C = 1) = o(E(X)^2)$ .

*Proof.* First let us estimate the number of ordered pairs  $(A, C)$  for which  $|A \cap C| < \sqrt{n}$ . Note that

$$\begin{aligned} \sum_{|A \cap C| < \sqrt{n}} 1 &= \binom{n}{a} \sum_{0 \leq i < \sqrt{n}} \binom{n-a}{a-i} \binom{a}{i} \leq \sqrt{n} \binom{n}{a} \binom{n-a}{a} \binom{a}{\sqrt{n}} \\ (3.6) \quad &= 2^{n(H(\alpha) + H(\frac{\alpha}{1-\alpha})(1-\alpha) + o(1))}. \end{aligned}$$

Now we will bound  $\Pr(X_A = X_C = 1)$  for fixed  $a$ -sets  $A$  and  $C$ . Let  $S \subset A \setminus C$  be a set of size  $s = |S| = \lfloor \sqrt{n} \rfloor$ . Define a new indicator random variable  $Y$  which takes the value 1 if and only if  $|B(C) \setminus S| \leq (\lambda_0 + \varepsilon)n - a$ . Clearly,  $X_C \leq Y$  and

$$\begin{aligned} \Pr(Y = 1) &= \Pr(Bin(n - a - s, p(a)) \leq (\lambda_0 + \varepsilon)n - a) \\ &\leq 2^{s+o(1)} \sum_{0 \leq i \leq (\lambda_0 + \varepsilon)n - a} \binom{n-a}{i} 2^{-(n-a)} = 2^{s+o(1)} \Pr(X_A = 1). \end{aligned}$$

Now if we condition on the existence or otherwise of all edges  $F'$  between  $C$  and  $V \setminus S$ , then if  $x \in V \setminus A$ ,

$$\Pr(x \in B(A) | F' \text{ and } F'') \in \left[ \frac{1 - (1 - 2p)^s}{2}, \frac{1 + (1 - 2p)^s}{2} \right],$$



where  $F''$  is the set of edges between  $x$  and  $A \setminus S$ . This implies that

$$\Pr(X_A = 1|Y = 1) = \sum_{0 \leq i \leq (\lambda_0 + \varepsilon)n - a} \binom{n - a}{i} 2^{-(n-a) + O(\sqrt{n})} = 2^{O(\sqrt{n})} \Pr(X_A = 1).$$

Consequently,

$$\Pr(X_A = X_C = 1) \leq \Pr(X_A = Y = 1) \leq 2^{O(\sqrt{n})} \Pr(X_A = 1)^2.$$

Hence, (3.6) implies

$$\sum_{|A \cap C| < \sqrt{n}} \Pr(X_A = X_C = 1) \leq 2^{n(H(\alpha) + H(\frac{\alpha}{1-\alpha})(1-\alpha) + o(1))} \Pr(X_A = 1)^2.$$

To complete the proof it is enough to note that

$$E(X)^2 = 2^{n(2H(\alpha) + o(1))} \Pr(X_A = 1)^2$$

and

$$2H(\alpha) > H(\alpha) + H\left(\frac{\alpha}{1-\alpha}\right)(1-\alpha).$$

Indeed, the last inequality follows from the strict concavity of the entropy function, since then  $(1-\alpha)H(\frac{\alpha}{1-\alpha}) + \alpha H(0) \leq H(\alpha)$  with the equality for  $\alpha = 0$  only.  $\square$

Now we show that  $f(G_{n,p}) \geq (\lambda_0 - \varepsilon)n$ . We show that

$$\sum_{1 \leq a \leq (\lambda_0 - \varepsilon)n} \binom{n}{a} \Pr(\text{Bin}(n - a, p(a)) \leq (\lambda_0 - \varepsilon)n - a) = o(1).$$

As in previous sections, we split this sum into two sums, but this time we make the break into  $1 \leq a \leq (\log n)^2$  and  $(\log n)^2 < a \leq (\lambda_0 - \varepsilon)n$ , respectively. In order to estimate the first sum we use the Chernoff bounds with deviation  $1 - \theta$  from the mean where

$$\theta = 1 - \frac{(\lambda_0 - \varepsilon)n - a}{(n - a)p(a)} \geq 1 - \frac{\lambda_0 - \varepsilon}{p(a)} \geq 1 - \frac{\lambda_0 - \varepsilon}{\lambda_0} = \frac{\varepsilon}{\lambda_0}.$$

Consequently,

$$\begin{aligned} \sum_{2 \leq a < (\log n)^2} \binom{n}{a} \Pr(\text{Bin}(n - a, p(a)) \leq (\lambda_0 - \varepsilon)n - a) \\ \leq (\log n)^2 \binom{n}{(\log n)^2} \exp\{-\Omega((\log n)^4)\} \leq \exp\{-\Omega((\log n)^4)\} = o(1). \end{aligned}$$

Now we bound the second sum corresponding to  $(\log n)^2 < a \leq (\lambda_0 - \varepsilon)n$ :

$$\begin{aligned} \sum_{(\log n)^2 \leq a \leq (\lambda_0 - \varepsilon)n} \binom{n}{a} \Pr(\text{Bin}(n - a, p(a)) \leq (\lambda_0 - \varepsilon)n - a) \\ = 2^{n(h(\lambda_0 - \varepsilon) + O(\log n/n))} = o(1). \end{aligned}$$

**4. General graphs.** Here we present the proof of Theorem 1.2. First, we prove a weaker result  $f(n) \leq (0.440\dots + o(1))n$ .

Suppose we aim at showing that  $f(n) \leq \lambda n$ . We fix some  $\alpha$  and  $\rho$  and let  $a = \alpha n$  and  $r = \rho n$ . For each  $a$ -set  $A$  let  $R(A)$  consist of all sets that have Hamming distance at most  $r$  from  $B(A)$ . If

$$(4.1) \quad \binom{n}{a} \sum_{i=0}^r \binom{n}{i} = 2^{n(H(\alpha)+H(\rho)+o(1))} > 2^n,$$

then there are  $A, A'$  such that  $R(A) \cap R(A') \ni C$  is nonempty. This means that  $C$  is within Hamming distance  $r$  from both  $B = B(A)$  and  $B' = B(A')$ . Thus  $|B \triangle B'| \leq 2r$ .

Let all vertices in  $A'' = A \triangle A'$  flip their neighbors, i.e., execute operation  $O_1$ . The only vertices outside of  $A''$  that can have an odd number of neighbors in  $A''$  are restricted to  $(B \triangle B') \cup (A \cap A')$ . Thus

$$(4.2) \quad f(G) \leq |A \triangle A'| + |(B \triangle B') \cup (A \cap A')| \leq 2a + 2r = 2n(\alpha + \rho).$$

Consequently, we try to minimize  $\alpha + \rho$  subject to  $H(\alpha) + H(\rho) > 1$ . Since the entropy function is strictly concave, the optimum satisfies  $\alpha = \rho$ ; otherwise replacing each of  $\alpha, \rho$  by  $(\alpha + \rho)/2$  we strictly increase  $H(\alpha) + H(\rho)$  without changing the sum. Hence, the optimum choice is

$$\alpha = \rho \approx 0.11002786443835959\dots,$$

the smaller root of  $H(x) = 1/2$ , proving that  $f(n) \leq (0.440\dots + o(1))n$ .

In order to obtain a better constant we modify the approach taken in (4.1). Let us take  $\delta = 0.275$ ,  $\alpha = 0.0535$ ,  $a = \lfloor \alpha n \rfloor$ ,  $d = \lfloor \delta n \rfloor$ . Look at the collection of sets  $B(A)$ ,  $A \in \binom{[n]}{a}$ . This gives  $\binom{n}{a} = 2^{n(H(\alpha)+o(1))}$  binary  $n$ -vectors.

We claim that some two of these vectors are at distance at most  $d$ . If not, then inequality (5.4.1) in [4] says that

$$H(\alpha) + o(1) \leq \min\{1 + g(u^2) - g(u^2 + 2\delta u + 2\delta) : 0 \leq u \leq 1 - 2\delta\},$$

where  $g(x) = H((1 - \sqrt{1-x})/2)$ . In particular, if we take  $u = 1 - 2\delta = 0.45$ , we get  $0.30108 + o(1) \leq 0.30103$ , a contradiction.

Thus, we can find two different  $a$ -sets  $A$  and  $A'$  such that  $|B(A) \triangle B(A')| \leq d$ . As in (4.2), we can conclude that  $f(G) \leq 2a + d \leq (0.382 + o(1))n$ .

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